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Age-gaps in sexual partnerships: seeing beyond ‘sugar daddies’

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Abstract

We examine for the first time age-mixing in sexual relationships in a population with very high HIV incidence and prevalence in rural South Africa. The highest levels of age assortativity (the pairing of like with like) were casual partnerships reported by men, the lowest levels were spousal relationships reported by women. Given the age–sex distribution of HIV prevalence in this population, interventions to decrease age-gaps in spousal relationships may be effective in reducing HIV incidence.

Age-mixing in sexual relationships may influence the spread of infectious diseases [1–4]. Information on age-mixing is crucial for investigating HIV epidemic dynamics [5–7] but rarely available for sub-Saharan Africa (SSA), with few notable exceptions [8]. In South Africa, age-gaps in sexual relationships have been reported as the proportion of partners in 5-year age categories around the age of the reporting partner [9,10]. However, no studies in South Africa have described continuously measured age-gaps, presented age-gap information by relationship type, or assessed the degree of age assortativity (the pairing of ‘like with like’ with respect to age) [11].

We used data from one of the largest HIV surveillances in SSA [12–14] in a community in rural South Africa with both high adult HIV prevalence (21%) [15] and incidence (3 per 100 person-years) [13]. Women and men aged 15 years or older were eligible to participate [16]. Participants were asked about their most recent sexual partners (up to three) and, for each partner, about the ‘relationship to that partner’ and whether the partner was ‘older, younger or about the same age’. A follow-up question ascertained the age difference in years, if same age was not reported [17]. Although it is possible that age self-reporting is biased, our data were collected within an established longitudinal demographic surveillance system in which age is ascertained at repeat visits, potentially reducing misreporting [18].

We calculated the age-gap in each relationship by subtracting the woman's age from the man's age. We constructed mixing matrices stratified by sex of respondent and relationship type with ages divided into 5-year categories. We quantified age assortativity by sex and relationship type by computing median age-gaps and Gupta *et al.*'s Q [19] on the mixing matrices:

$$Q = \frac{\sum_{i=1}^N w_i - 1}{N - 1}$$

in which w_i are the eigenvalues of the mixing matrix and N is the number of age categories. Q quantifies the proportion of partnerships that occur within the same age group [20]. Q can be used to compare different samples of relationships, as long as the age categories of the mixing matrices are the same. Q is 1 when all relationships occur between people in the same age group, 0 when age group mixing is proportional and $-1/(N-1)$ when all relationships occur between people in different age groups. Although Q does not have a natural interpretation outside these three extreme values [20], it does allow ranking of relationships on a continuum from maximum disassortativity to maximum assortativity. To estimate the variance of Q , we carried out 10 000 bootstrap replications with replacement [21]. We fit multivariable models to identify the shape of the association between age-gap and independent variables (age, sex and relationship type), using likelihood-ratio tests to compare nested models.

One thousand three hundred and forty-nine men and 2768 women provided information on a total of 4437 partners. Only 152 men and 96 women provided information on two partners and only 31 men and five women on three partners. See Supplemental Digital Content 1, <http://links.lww.com/QAD/A116>, for mixing matrices and Q statistics and Supplemental Digital Content 2, <http://links.lww.com/QAD/A117>, for the distributions of age-gap by sex and partnership type. Only 5% of men were younger than their female partners. All relationships were age-assortative (i.e. $Q > 0$). Spousal relationships (which comprised 15% of relationships) had the lowest levels of age assortativity in men and women. Relationships reported between casual partners had higher Q than spousal relationships when reported by women (0.357 vs. 0.202, $P < 0.001$) or men (0.393 vs. 0.265, $P = 0.09$).

After comparing every combination of main-effect models and models with every combination of two-way and three-way interactions, the best fitting model included all two-way interaction terms between age (modelled quadratically), sex and relationship type and their main effects (see Fig. 1 for the age-gaps predicted from the model). Among male respondents, age-gaps increased for all four relationships types through age 50 and decreased after age 50 (in casual relationships), after the mid-60s (spousal relationships) and after the mid-70s (former spousal and casual relationships). Age-gaps between women and their causal partners increased until the early 30s, after which they decreased rapidly. Age-gaps between women and their former spouses or partners held fairly constant until age 35 and then decreased. The age-gap between female respondents and their current partners reached a maximum of about 4.5 years in women above 40 years of age.

Past studies have linked age-gaps to increased risk of HIV infection in SSA [1,7,22–25]. Our study provides a finely detailed description of the age-mixing patterns in a rural population in SSA, in which HIV prevalence peaks 5 years earlier in women than in men [12,26]. When a young woman enters a sexual relationship with an older man, as frequently occurs, we would thus expect that she is at a higher risk of contracting HIV than if she entered a relationship with a man of her own age. Age-gaps are often hypothesized to be largest in casual relationships, coinciding with compensation for the younger female partner, for example gifts from a ‘sugar daddy’ (a man at least 10 years older than his casual partner) [27]. We find that regardless of the sex or age of the respondent, spousal relationships have substantially larger age-gaps than casual relationships and ‘sugar daddy’ relationships are quite rare in this population. Age-gaps in spousal relationships may be large because men

may need to save money over a long time to accumulate enough wealth to pay *lobola*, the traditional bride price in this community [2,28]. By the time a man can afford *lobola*, he may have engaged in several relationships with casual partners, increasing his and his future wife's risk of HIV infection.

Future studies should estimate the effect of age-gaps on the hazard of HIV acquisition and try to assess whether any such effects are due to the HIV age distributions in men and women or due to differences in sexual risk-taking by age-gaps [22]. If age-gaps in sexual relationships are indeed important factors in HIV acquisition, interventions that reduce the gaps should be developed [29,30], for example informing women that older men are more likely to be HIV-infected than younger men or providing loans for *lobola* for young men.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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M.Q.O. and T.B. jointly planned and performed the analyses, and wrote and edited the article. M.-L.N., F.T. and M.L. contributed at all stages of the analysis plans and edited the article for substantive content. All authors approved the final manuscript.

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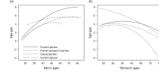


Fig. 1. Age-gap by relationship type
(a) In male respondents. (b) In female respondents.